The 2015 United Nations/Japan Space Weather Workshop

Spatio-temporal characteristics of the equatorial ionization anomaly (EIA) in the East African sector using ionospheric tomography

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Venue: Kyushu University, Fukuoka, Japan



# Outline





- 3 Data and methods of analysis
  - Presults and discussions







# A glance on Bahir Dar University!





### **Project overview**



### $Cont \cdots$

Theoretical

Solve conservation equations numerically as a function of space and time

Require magnetospheric and atmospheric inputs

Outputs: Calculate plasma density, temperature and flow velocity

Are powerful to understand the physical and chemical processes of the upper atmosphere

Example: The Sheffield Coupled Thermosphere Ionosphere Plasmasphere Model (CTIP)





### $Cont \cdots$

Empirical

Analytical descriptions of the ionosphere with functions derived from experimental data or adopted from physical models

The two known empirical models are International reference ionosphere (IRI) and NeQuick

#### IRI



#### Its data sources include

- Coefficients such as foF2 from ground-based sounders
- Incoherent scatter measurements
- Topside sounder ionograms
- In-situ measurements



### $\operatorname{Cont}\cdots$



# Outputs

- Electron temperature
- Ion temperature
- Total electron content (TEC)





### $\operatorname{Cont}\cdots$





### $\mathbf{Cont}\cdots$

Can measure TEC

Installing various GPS receivers around the world can make this measurement

It is possible to create a global view of the ionosphere by utilizing all possible data merging them together and interpolating them

The TEC can be inverted to estimate the electron density at different height (Ionospheric tomography)









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Kassa (WaGRL, BDU, Ethiopia) Spatiotemporal characteristics 02-06 March 2015

### $Cont \cdots$





Figure: Distribution of irregularities from year 1999-2004 over Africa (Groves et al., 2012).



### **Objectives**







#### lonosphere in the nutshell

### **Ionospheric regions**



#### Ionosphere in the nutshell

### Formation and dynamics of the EIA



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Spatiotemporal characteristics

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### Data







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Spatiotemporal characteristics





### GPS signal and atmospheric effects



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**Spatiotemporal characteristics** 

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### **Ionospheric delay**

- When passing through the atmosphere, the electromagnetic signals experience changes in velocity (speed and direction) due to the refraction.
- According to Fermat's principle, the measured range l is given by the integral of the refractive index n along the ray path from the satellite to receiver:

$$I = \int_{\rm rp} n dI. \tag{1}$$

Accordingly, the signal's delay can be written as

$$\Delta = \int_{\rm rp} ndl - \int_{\rm s1} dl, \qquad (2)$$

where the second integral is the Euclidean distance between the satellite and receiver.

- Eq. (2) includes both the signal bending and propagation delay.
- One of the possible simplification can be obtained by approximating the first integral along the straight line between the satellite and receiver:

$$\Delta = \int_{\mathrm{sl}} (n-1) dl. \tag{3}$$





### Dual frequency GPS signal in extracting ionospheric TEC

- A medium where the wave propagation speed and hence the refractive index depend on the frequency or the angular frequency ω and the wave number k are not proportional is a dispersive medium.
- This is also valid for the ionosphere where  $\omega$  and k are related, to a first approximation

$$\omega^2 = c^2 k^2 + \omega_p^2, \tag{4}$$

where c is the propagation speed of a signal in a vacuum and  $\omega_p = 2\pi f_p$ , where  $f_p = 8.98\sqrt{N_e}$  in Hz and  $N_e$  is the electron density (in  $\frac{e}{m^3}$ ). Eq. (4) is called the relation of dispersion of the ionosphere, and  $\omega_p$  is called the critical frequency of the ionospheric plasma.

Four GPS range measurements

$$\begin{split} P_{i} &= \rho + a_{gr,i} + \tau_{i}^{r} + \tau_{i}^{*} \\ P_{2} &= \rho + a_{gr,i} + \tau_{i}^{r} + \tau_{i}^{*} \\ P_{1} &= \rho + a_{gr,i} + \tau_{i}^{r} + \tau_{i}^{*} \\ L_{1} &= \rho + a_{gr,i} + N_{1}\lambda_{1} + \varepsilon_{i}^{r} + \varepsilon_{i}^{*} \\ L_{2} &= \rho + a_{gr,i} + N_{2}\lambda_{2} + \varepsilon_{2}^{r} + \varepsilon_{2}^{*} \\ \text{where } P_{i} &= \text{measured ode ranges} \\ L_{i} &= \text{measured phase ranges} (i = 1, 2) \text{ on the} \\ L_{i} \text{ and } L_{2} \text{ frequencies} \\ a_{gr,i} &= \text{group ionospheric delay term} \\ a_{gr,i} &= \text{phase ionospheric delay term} \\ \lambda_{i} &= \text{carrier wavelength} \\ N_{i} &= \text{int eger cycle ambiguity} \\ \tau_{i}, \varepsilon_{i} &= \text{dispersion components of} \end{split}$$

the satellite and the reciever hardware delays

Usefull relations

$$n_{ph} = 1 - \frac{40.3}{f^2} Ne$$
  
 $n_{gr} = 1 + \frac{40.3}{f^2} Ne$ 

$$a_{phJ} = -\frac{40.3}{f^2} \int \text{Nedl} = -\frac{40.3}{f^2} \text{I}$$

$$a_{prJ} = \frac{40.3}{f^2} \int \text{Nedl} = \frac{40.3}{f^2} \text{I}, \text{ where } \text{I} = \text{sTEC} = \int \text{Nedl}$$

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**Spatiotemporal characteristics** 

### Cont···

If we neglect terms other than \(\rho\), \(\Delta\)<sub>ph</sub>, i and \(\Delta\)<sub>gr</sub>, i we may have

PI

$$P_1 = \rho + \Delta_{gr,1}$$
(5)

$$P_2 = \rho + \Delta_{gr,2} \tag{6}$$

$$L_1 = \rho + \Delta_{ph,1} \tag{7}$$

$$L_2 = \rho + \Delta_{ph,2}. \qquad (8)$$

(11)

The dual frequency differencing of the code and phase ranges gives the ionospheric combinations  $P_I$  and  $L_I$ : ۰

$$P_{I} = P_{2} - P_{1} = \rho + \Delta_{gr,2} - (\rho + \Delta_{gr,1}) = \rho + \frac{f_{1}^{2}}{f_{2}^{2}} \Delta_{gr,1} - \rho - \Delta_{gr,1}$$
(9)

$$L_{I} = L_{1} - L_{2} = \rho + \Delta_{ph,1} - (\rho + \Delta_{ph,2}) = \rho + \Delta_{ph,1} - \rho - \frac{f_{1}^{2}}{f_{2}^{2}} \Delta_{ph,1}.$$
 (10)

Thus we have

$$P_{I} = P_{2} - P_{1} = \frac{f_{1}^{2}}{f_{2}^{2}} \Delta_{gr,1} - \Delta_{gr,1} = \Delta_{gr,1} \left(\frac{f_{1}^{2}}{f_{2}^{2}} - 1\right)$$
$$L_{I} = L_{1} - L_{2} = \Delta_{ph,1} - \frac{f_{1}^{2}}{f_{2}^{2}} \Delta_{ph,1} = \Delta_{ph,1} \left(1 - \frac{f_{1}^{2}}{f_{2}^{2}}\right).$$



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• By using the relations 
$$\Delta_{ph,1} = -\frac{40.3}{f_1^2} sTEC_P$$
 and  $\Delta_{gr,1} = \frac{40.3}{f_1^2} sTEC_L$  we get

$$P_2 - P_1 = \frac{40.3}{f_1^2} sTEC\left(\frac{f_1^2}{f_2^2} - 1\right) = 40.3\left(\frac{f_1^2 - f_2^2}{f_1^2 f_2^2}\right) sTEC_P$$
(13)

$$L_{1} - L_{2} = -\frac{40.3}{f_{1}^{2}} sTEC\left(1 - \frac{f_{1}^{2}}{f_{2}^{2}}\right) = -40.3 \left(\frac{f_{2}^{2} - f_{1}^{2}}{f_{1}^{2}f_{2}^{2}}\right) sTEC_{L}.$$
 (14)

From Eqs. (13) and (14) we get

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$$sTEC_P = \frac{1}{40.3} \left( \frac{f_1^2 f_2^2}{f_1^2 - f_2^2} \right) (P_2 - P_1)$$
(15)

$$sTEC_L = \frac{1}{40.3} \left( \frac{f_1^2 f_2^2}{f_1^2 - f_2^2} \right) (L_1 - L_2), \tag{16}$$

where  $sTEC_P$  is the slant total election content from code measurements and  $sTEC_L$  is the slant total election content from phase measurements.

If we consider the differencing of the code and the phase ranges including the biases and other terms we get

$$P_{I} = P_{2} - P_{1} = 40.3 \left(\frac{1}{f_{2}^{2}} - \frac{1}{f_{1}^{2}}\right) sTEC_{P} + B_{r} + B_{s}$$
 (17)

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- The code range ionospheric combination P<sub>1</sub> is absolute, but noisy.
- On the other hand, the phase combination, L<sub>1</sub>, is smooth and can provide very accurate measurements of TEC changes, but is offset from the code range combination by a bias related to the unknown integer cycle ambiguity, M<sub>1</sub> Å<sub>1</sub> − M<sub>2</sub>Å<sub>2</sub>.
- The ambiguity (phase leveling) is removed by averaging (sTEC<sub>P</sub> sTEC<sub>L</sub>) over a satellite pass (phase connecting arc) as

$$sTEC_{leveled} = sTEC_L - mean(sTEC_L - sTEC_P)$$

(19)

This "levels" the TEC to the unambiguous.





# Methods of analysis





### Formulating statistical linear inverse problem

The sTEC and the ionospheric electron density (Ne) can be related by

$$sTEC = \int_{I} Ne(\phi, \lambda, h, t) ds, \qquad (20)$$

where  $\phi$ ,  $\lambda$ , h and t stand for latitude, longitude, altitude and time variables respectively and l refers the ray propagation path of each GPS satellite-receiver pair.

For 2D case, the sTEC can be formulated as

$$sTEC = \int_{I} Ne(\phi, h) ds.$$
 (21)

The discretized version of Eq. (21) becomes

$$\mathsf{sTEC}_{i} = \sum_{j=1}^{n} A_{ij} \mathsf{Ne}_{j} + \varepsilon_{i}, \tag{22}$$



where *i* is the *i*<sup>th</sup> measurement along propagation path of each GPS satellite-receiver pair.

### $\operatorname{Cont}\cdots$

Eq. (22) can be generally written in simple matrix notation as:

$$\mathbf{y}_{m \times \mathbf{1}} = \mathbf{A}_{m \times n} \mathbf{x}_{n \times \mathbf{1}} + \varepsilon_{m \times \mathbf{1}}, \tag{23}$$

where *n* is the number of pixels in the image, *m* is the number of *sTEC* measurements, *y* is a column vector of the *m* known *sTEC* measurements, A is an  $m \times n$  projection matrix with  $A_{ij}$  being the length of ray *i* traversing pixel *j*, *x* is a column vector consisting of all the unknown electron densities (Ne) in all the pixels, and *c* is a column vector associated with additive measurement noises.

Let the electron density vector, x and the measurement error vector, ε be independent, jointly distributed random variables with x ~ Normal(μ<sub>n</sub>, Σ<sub>r</sub>) and ε ~ Normal(0<sub>m</sub>, Σ<sub>ε</sub>) where μ<sub>n</sub>, Σ<sub>r</sub>, 0<sub>m</sub> and Σ<sub>ε</sub> are mean of the unknown, covariance of the unknown, covariance of the unknown, mean and covariance of the measurement error respectively.

If A be an m × n matrix, we can define a new random vector

$$\mathbf{y} = \mathbf{A}\mathbf{x} + \varepsilon. \tag{24}$$





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### Solving statistical linear inverse problem in Bayesian framework

If the measurement errors are Gaussian with zero mean, the a posteriori probability density of the unknown, x, according to Bayesian framework, given its priori density, π<sub>pr</sub>(x), and measurement vector, y, can be spelled out as

$$\pi_{\text{post}}(\mathbf{x}) = \pi(\mathbf{x}|\mathbf{y}) = \frac{\pi_{\text{pr}}(\mathbf{x}) \times \pi(\mathbf{y}|\mathbf{x})}{\pi(\mathbf{y})},$$
(25)

where  $\pi_{\mathrm{pr}}(x)$  is the joint a priori density of the unknowns and the marginal density  $\pi(y) = \int_{\mathbb{R}} m \pi(x, y) dx = \int_{\mathbb{R}} m \pi(y|x) \pi_{\mathrm{pr}}(x) dx$  is can be seen as a norming constant and usually of little importance.

- The expression  $\pi(y|x) \propto \exp\left[-\frac{1}{2}(y Ax)^T \Sigma_{\varepsilon}^{-1}(y Ax)\right]$  is called the likelihood function, as it expresses the likelihood of different measurement outcomes when x = x is given. Maximum Likelihood Estimate (MLE)
- If we don't consider a prior information,  $\pi_{pr}(x)$  can be constant and thus the most probable values of the unknown can be computed by maximizing  $\pi(x|y) = \pi(y|x) \propto \exp\left[-\frac{1}{2}(y Ax)^T \Sigma_{\varepsilon}^{-1}(y Ax)\right]$  with respect to the unknown x. That means differentiating the above expression with respect to x and letting to be zero.

$$-\left(\frac{d}{dx}(\mathbf{A}x)\right)^{T}\Sigma_{\varepsilon}^{-1}(\mathbf{y}-\mathbf{A}x)-(\mathbf{y}-\mathbf{A}x)^{T}\Sigma_{\varepsilon}^{-1}\frac{d}{dx}(\mathbf{A}x)=0.$$
(26)



### $\operatorname{Cont}\cdots$

- This implies that the matrix multiplication,  $A^T \Sigma_{\varepsilon}^{-1} B = B^T \Sigma_{\varepsilon}^{-1} A$ , is valid.
- Applying the above matrix multiplication identity into Eq. (26) gives

$$-\left(\frac{d}{dx}(\mathbf{A}\mathbf{x})\right)^{T}\Sigma_{\varepsilon}^{-1}(\mathbf{y}-\mathbf{A}\mathbf{x})-\left(\frac{d}{dx}(\mathbf{A}\mathbf{x})\right)^{T}\Sigma_{\varepsilon}^{-1}(\mathbf{y}-\mathbf{A}\mathbf{x})=0.$$
 (27)

This in turn yields

$$-2\left(\frac{d}{dx}(\mathbf{A}\mathbf{x})\right)^{\mathsf{T}}\Sigma_{\varepsilon}^{-1}(\mathbf{y}-\mathbf{A}\mathbf{x})=0. \tag{28}$$

Differentiation and further mathematical manipulations simplify the above equation into

$$\mathbf{A}^{\mathsf{T}} \boldsymbol{\Sigma}_{\varepsilon}^{-1} (\mathbf{y} - \mathbf{A}\mathbf{x}) = \mathbf{0}$$
$$\mathbf{A}^{\mathsf{T}} \boldsymbol{\Sigma}_{\varepsilon}^{-1} \mathbf{y} = \mathbf{A}^{\mathsf{T}} \boldsymbol{\Sigma}_{\varepsilon}^{-1} \mathbf{A}\mathbf{x}.$$
 (29)

#### The resulting solution, which is known as the Maximum Likelihood Estimate (MLE), is given by

$$\mathbf{x}_{\mathrm{MLE}} = (\mathbf{A}^{\mathrm{T}} \boldsymbol{\Sigma}_{\varepsilon}^{-1} \mathbf{A})^{-1} \mathbf{A}^{\mathrm{T}} \boldsymbol{\Sigma}_{\varepsilon}^{-1} \mathbf{y}.$$





# Maximum A Posteriori (MAP) Estimate

- An obvious limitation occurs on the maximum likelihood estimate of Eq. (30) when the inverse matrix  $(A^T \Sigma_c^{-1} A)^{-1}$  do not exist.
- However, even when the inverse matrix exists, it can still be ill-conditioned (become nearly singular).
- In this case our solution would be extremely unstable and unrealistic.
- The instability of Eq. (30) exhibit too strong point-to-point variations in the inversion.
- To overcome this problem, we need to find a way of limiting these variations (we have to apply regularization methods).
- To do so, let us consider two neighboring grid points (pixels) p<sub>i</sub> and p<sub>j</sub> located at the same altitude or at the same latitude. If we assume that the difference of electron densities at these points is close to zero, we incorporate some error e<sup>ij</sup>.
- This can be formulated as

$$\mathbf{0} = \mathbf{x}_{i} - \mathbf{x}_{j} + \varepsilon_{r}^{ij}. \tag{31}$$

If we do this for all vertical and horizontal differences of neighboring grid points, the resulting system of equations can be expressed in a single matrix form as

$$\mathbf{0} = \mathbf{A}_{\mathbf{r}}\mathbf{x} + \varepsilon_{\mathbf{r}},\tag{32}$$

where  $A_r$  is regularization matrix with elements +1, -1 and 0 in appropriate places. We can now combine Eq. (24) and (32) to get a single matrix equation as

 $\begin{pmatrix} \mathbf{y} \\ \mathbf{0} \end{pmatrix} = \begin{pmatrix} \mathbf{A} \\ \mathbf{Ar} \end{pmatrix} \mathbf{x} + \begin{pmatrix} \varepsilon \\ \varepsilon \mathbf{r} \end{pmatrix}.$ 



### $Cont \cdots$

• The regularized a posterior density,  $\pi_{post}(x) = \pi(x|y, 0)$ , can be written as

$$\pi_{\mathrm{post}}(\mathbf{x}) = \pi(\mathbf{x}|\mathbf{y},\mathbf{0}) \quad \propto \quad \exp\left[-\frac{1}{2}(\mathbf{A}_{\mathrm{r}}\mathbf{x})^{\mathrm{T}}\boldsymbol{\Sigma}_{\mathrm{r}}^{-1}(\mathbf{A}_{\mathrm{r}}\mathbf{x})\right] \times \exp\left[-\frac{1}{2}(\mathbf{y}-\mathbf{A}\mathbf{x})^{\mathrm{T}}\boldsymbol{\Sigma}_{\varepsilon}^{-1}(\mathbf{y}-\mathbf{A}\mathbf{x})\right],$$

where  $\pi_{\mathrm{pr}}(x) \propto \exp\left[-\frac{1}{2}(A_r x)^{\mathrm{T}} \Sigma_r^{-1}(A_r x)\right]$  is the distribution of difference prior with covariance matrix  $\Sigma_r$ .

 By applying a similar maximization for the posteriori distribution given by Eq. (34), we can compute the most probable values of the unknowns as

$$\left(\frac{d}{dx}(\mathbf{A}_{\mathbf{r}}\mathbf{x})\right)^{\mathsf{T}}\Sigma_{\mathbf{r}}^{-1}(\mathbf{A}_{\mathbf{r}}\mathbf{x}) - \left(\frac{d}{dx}(\mathbf{A}\mathbf{x})\right)^{\mathsf{T}}\Sigma_{\varepsilon}^{-1}(\mathbf{y}-\mathbf{A}\mathbf{x}) - (\mathbf{y}-\mathbf{A}\mathbf{x})^{\mathsf{T}}\Sigma_{\varepsilon}^{-1}\frac{d}{dx}(\mathbf{A}\mathbf{x}) = 0.$$
(34)

Handling the differentiation with respect to x and applying the matrix multiplication identity,  $A^T \Sigma_{\varepsilon}^{-1} B = B^T C_{\varepsilon}^{-1} A$ , will lead us to have

$$2\mathbf{A}_{r}^{T}\boldsymbol{\Sigma}_{r}^{-1}\mathbf{A}_{r}\mathbf{x} - 2\mathbf{A}^{T}\boldsymbol{\Sigma}_{\varepsilon}^{-1}\mathbf{y} + 2\mathbf{A}^{T}\boldsymbol{\Sigma}_{\varepsilon}^{-1}\mathbf{A}\mathbf{x} = 0.$$
(35)

Further rearrangements result in

$$(\mathbf{A}_{r}^{T}\boldsymbol{\Sigma}_{r}^{-1}\mathbf{A}_{r}+\mathbf{A}^{T}\boldsymbol{\Sigma}_{\varepsilon}^{-1}\mathbf{A})\mathbf{x}-\mathbf{A}^{T}\boldsymbol{\Sigma}_{\varepsilon}^{-1}\mathbf{y}=\mathbf{0}. \tag{36}$$

Solving for x = x<sub>MAP</sub> gives

$$\mathbf{x}_{\mathrm{MAP}} = (\mathbf{A}^{\mathrm{T}} \boldsymbol{\Sigma}_{\varepsilon}^{-1} \mathbf{A} + \mathbf{A}_{\mathrm{r}}^{\mathrm{T}} \boldsymbol{\Sigma}_{\mathrm{r}}^{-1} \mathbf{A}_{\mathrm{r}})^{-1} \mathbf{A}^{\mathrm{T}} \boldsymbol{\Sigma}_{\varepsilon}^{-1} \mathbf{y}.$$



Eq. (37) is commonly known as Maximum of a Posteriori Estimate (MAP) of the unknowns.

### $\mathbf{Cont}\cdots$

- ullet We have used Chapman profile as a priori profile of the Ne in the form of priori covariance matrix,  $\Sigma_{
  m r}.$
- The variance (error) estimate of the statistical linear inversion can be obtained from the posterior covariance matrix, pcovmat = (A<sup>T</sup>Σ<sub>r</sub><sup>-1</sup>A + A<sup>T</sup><sub>r</sub>Σ<sub>r</sub><sup>-1</sup>A<sub>r</sub>)<sup>-1</sup>, in such a way that the diagonal elements of this matrix belong to the square of the variances of the estimated error.
- Thus the variances, which are the square root of the diagonal elements of the posterior covariance matrix, can be considered as the error estimate of the employed inversion approach.



### Sample ray path and the EIA movement/dimension from vTEC



# Validation







# Diurnal and seasonal variations of the EIA



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# Daily variations of the EIA





# $\operatorname{Cont}\cdots$





### Storm-time responses: 09 March 2012 storm





### Temporal variations of NmF2 of the EIA









### Spatial variations of the EIA

- The spatial variations of the EIA can be seen in terms of its altitude and latitude profile of the EIA at 21:00 LT in a particular day (DOY 356) and season (December solstice).
- This can be seen from the middle panel of Fig. (35) at 21:00 LT. In this Fig., it is manifested that the maximum of the EIA is appeared around 312 km and its thickness along the altitude is found to be between 180 km to 450 km (270 km) during December solstice.
- Moreover, the latitude profile of the EIA advocated that the ionosphere is pinched around the equator and thickened around  $6.3^{\circ}$  N to  $11^{\circ}$  N GeoMLat (with a thickness of  $5^{\circ} \simeq 555$  km).
- This might be explained by an ionization transport mechanism commonly referred to as the equatorial fountain effect and the later downward diffusion of ionization along the magnetic field lines (Andreeva et al., 2000).



#### Conclusions

## Conclusions

- We have produced the images of the EIA by applying Bayesian inversion theory.
- We have described the temporal, spatial and storm-time variations of the EIA. The diurnal variations of EIA was found to be dependent on the position of the sun over the horizon and the daily variations have been explained by the occurrence of storm events.
- We have also shown that the EIA behaves differently on different seasons and pronounced EIA has been observed during Equinox than solstice seasons.
- Even if a sharp rise in EIA from about 09:00 to 16:00 LT was observed at all seasons, the strength is high in equinoctial months followed by December and June solstices.
- Seasonal variations of the EIA depend not only on production and loss of ionization but also on the transport of plasma through winds and on thermospheric neutral composition variations.
- The strong EIA in equinox months are due to high values of solar intensity.
- Due to the combined effect of the trans-equatorial wind and the auroral equator-ward wind there is an occurrence of an asymmetry of the two crests of the EIA.



#### Conclusions

### $\operatorname{Cont}\cdots$

- During solstice the two crests move equator ward but during equinox the two crests move slightly poleward.
- The investigation on the spatial parameters has shown that the EIA varies along both altitude and latitude coordinates.
- The synthesis on the peak of the EIA (NmF2) has demonstrated a significant dependence of NmF2 on local time, day and season of the year.
- The storm-time analysis has also reveled that the EIA responds differently during geomagnetically disturbed and stable situations.
- The enhanced EIA at 18:00 UT (longer time after the onset of the storm) might be observed by the disturbance dynamo electric field due to the storm-time circulation (Fejer and Scherliess, 1995).
- In terms of the the appearance of the daytime EIA (the peak and the width), we have observed that the northern crest predominately appeared well whereas the southern anomaly relatively inhibited during the geomagnetic storm we considered.
- This implies that the effect of the storm has a clear latitudinal (spatial) dependance or means that the effects of storm has propagated along the latitude.



### Articles

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- Kassa T., Damtie B., Bires A., Yizengaw E. & Cilliers P. Spatio-temporal characteristics of the Equatorial Ionization Anomaly (EIA) in the East African region via ionospheric tomography during the year 2012. United Nations/Japan Workshop on Space Weather "Science and Data Products from ISWI Instruments", Fukuoka, Japan, 2-6 March 2015.
- 2 Aara'L Y. and Kassa T. Statistics of ionospheric amplitude scintillation and its correlation with phase scintillation over Bahir Dar using GPS SCINDA during solar maximum phase. Emerging Researchers National Conference in Science, Technology, Engineering and Mathematics (STEM), February 19-21 2015, Washington DC, USA.
  - Kassa T., Tessema F., Bires A. and Nigussie M. Ionospheric response to Annular solar eclipse of 15 January 2010 and 03 November 2013 in the Ethiopian zone, the 4<sup>th</sup> Gondar School of Science and Technology, January 2014, Gondar University, Ethiopia.
- Kassa T. Space Science activities and facilities at Washera Geospace and Radar Science Laboratory, Bahir Dar, Ethiopia, the 4<sup>th</sup> Gondar School of Science and Technology, January 2014, Gondar University, Ethiopia.
- **Kassa T.** Characterizing the Ethiopian Ionosphere via GPS Observables and 2D electron density profile, ISWI/SCOSTEP, 21 October to 01 November 2013, Nairobi, Kenya.
- **Kassa T.** Space Science activities at Washera Geospace and Radar Science Laboratory, ISWI/SCOSTEP, 21 October to 01 November 2013, Nairobi, Kenya.
- **Wassa T.** Characterizing Equatorial Ionization Anomaly in Ethiopian sector, Beacon Satellite symposium, Bath University, 08-12 July 2013, UK.
- 8 Kassa T. Effects of 07 April 2011 and 22 January 2012 geomagnetic storms on GPS ionospheric observables at Bahir Dar , the 1<sup>st</sup> Annual Science Conference (ASC) organized by Science College, BDU, 31 May-01 June 2013, Gondar University, Ethiopia.
  - **Kassa T.** Characterizing Equatorial Ionization Anomaly in Ethiopian sector, Ionospheric monitoring workshop, Hermanus, 24-25 January 2013, South Africa.

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